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Search for resonances decaying to $\eta_c \pi^+ \pi^-$ in two-photon interactions

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We report a study of the process $\gamma\gamma \rightarrow X \rightarrow \eta_c \pi^+ \pi^-$, where X stands for one of the resonances $\chi_{c2}(1P)$, $\eta_c(2S)$, $X(3872)$, $X(3915)$, or $\chi_{c2}(2P)$. The analysis is performed with a data sample of 473.9 fb^{-1} collected with the *BABAR* detector at the PEP-II asymmetric-energy electron-positron collider. We do not observe a significant signal for any channel, and calculate 90% confidence-level upper limits on the products of branching fractions and two-photon widths $\Gamma_{X \rightarrow \gamma\gamma} \mathcal{B}(X \rightarrow \eta_c \pi^+ \pi^-)$: 15.7 eV for $\chi_{c2}(1P)$, 133 eV for $\eta_c(2S)$, 11.1 eV for $X(3872)$ (assuming it to be a spin-2 state), 16 eV for $X(3915)$ (assuming it to be a spin-2 state), and 18 eV for $\chi_{c2}(2P)$. We also report upper limits on the ratios of branching fractions $\mathcal{B}(\eta_c(2S) \rightarrow \eta_c \pi^+ \pi^-) / \mathcal{B}(\eta_c(2S) \rightarrow K_S^0 K^+ \pi^-) < 10.0$ and $\mathcal{B}(\chi_{c2}(1P) \rightarrow \eta_c \pi^+ \pi^-) / \mathcal{B}(\chi_{c2}(1P) \rightarrow K_S^0 K^+ \pi^-) < 32.9$ at the 90% confidence level.

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Two-photon fusion events provide a useful production mode to study charmonium states with quantum numbers $J^{PC} = 0^{\pm+}, 2^{\pm+}, 4^{\pm+}, \dots, 3^{++}, 5^{++}, \dots$ [1,2]. Di-pion transitions among these states have been searched for only in the case of $\eta_c(2S) \rightarrow \eta_c \pi^+ \pi^-$ [3], in contrast to the narrower vector states, where di-pion transitions have been studied extensively. In particular, the transition amplitude for $\eta_c(2S) \rightarrow \eta_c \pi^+ \pi^-$ [4] is expected [5] to have

the same approximately linear dependence on the squared invariant mass of the di-pion system as the $\psi(2S) \rightarrow J/\psi \pi^+ \pi^-$ decay [6]. Phase-space integration of the squared amplitude, evaluated for the peak masses M_{η_c} and $M_{\eta_c(2S)}$ [1] of the η_c and $\eta_c(2S)$, respectively, yields $\Gamma(\eta_c(2S) \rightarrow \eta_c \pi^+ \pi^-) / \Gamma(\psi(2S) \rightarrow J/\psi \pi^+ \pi^-) \approx 2.9$. This leads to the branching fraction prediction $\mathcal{B}(\eta_c(2S) \rightarrow \eta_c \pi^+ \pi^-) = (2.2_{-0.6}^{+1.6})\%$, where the uncertainty is due to the uncertainty on the width of the $\eta_c(2S)$ [1]. This decay may be further suppressed due to the contribution of the chromomagnetic interaction to the decay amplitude [7].

In recent years, experiments have reported evidence for charmoniumlike states, such as the $X(3872)$ [8] and $Y(4260)$ [9,10], which do not fit well into the conventional $c\bar{c}$ picture. This has prompted much theoretical activity and proposals for new models [11]. Several studies of these states have been performed with the $J/\psi \pi^+ \pi^-$ final state [9,12], but no search using the $\eta_c \pi^+ \pi^-$ final state has been conducted. Such a search may shed light on the quantum numbers or the internal dynamics of these states. In particular, it has been suggested [13] that if the $X(3872)$ is the 1^1D_2 state η_{c2} , then the branching fraction $\mathcal{B}(X(3872) \rightarrow \eta_c \pi^+ \pi^-)$ could be significantly larger than $\mathcal{B}(X(3872) \rightarrow J/\psi \pi^+ \pi^-)$. The quantum numbers $J^{PC} = 2^{-+}$ of the η_{c2} are consistent with the results of an angular analysis of $X(3872) \rightarrow J/\psi \pi^+ \pi^-$ [14] and would allow production of $X(3872)$ in two-photon fusion.

We present herein a study of the process $\gamma\gamma \rightarrow X \rightarrow \eta_c \pi^+ \pi^-$, where X is one of the resonances $\chi_{c2}(1P)$, $\eta_c(2S)$, $X(3872)$, $X(3915)$, or $\chi_{c2}(2P)$, and the η_c is reconstructed in the final state $K_S^0 K^+ \pi^-$ [15].

The data sample was collected with the *BABAR* detector at the PEP-II asymmetric-energy e^+e^- collider located at the SLAC National Accelerator Laboratory. It consists of $429.1 \pm 1.9 \text{ fb}^{-1}$ collected at the energy of the $Y(4S)$ resonance, constituting the entire *BABAR* $Y(4S)$ data set, and $44.8 \pm 0.2 \text{ fb}^{-1}$ collected about 40 MeV below the $Y(4S)$ resonance. The *BABAR* detector is described in detail in Ref. [16].

Samples of Monte Carlo (MC) simulated events are analyzed with the same reconstruction and analysis procedures as the data sample, following a GEANT4-based [17] detector simulation [16]. Simulated background samples include $e^+e^- \rightarrow q\bar{q}$ continuum events ($q = u, d, s, c$) generated with JETSET [18], $Y(4S) \rightarrow B\bar{B}$ decays generated with EVTGEN [19] and JETSET, and $e^+e^- \rightarrow \tau^+\tau^-$ events generated with KK [20]. In order to study initial-state-radiation (ISR) background and the invariant-mass resolution, a sample of $e^+e^- \rightarrow \gamma\psi(2S)$ events with $\psi(2S) \rightarrow J/\psi \pi^+ \pi^-$ and $J/\psi \rightarrow K_S^0 K^+ \pi^-$ is generated with EVTGEN. The GAMGAM [21] generator is used to generate signal event samples for each of the X states studied, with the decay $X \rightarrow \eta_c \pi^+ \pi^-$ simulated with an amplitude that is uniform throughout the decay phase space, independent

of the final-state kinematic variables. The decay $\eta_c \rightarrow K_S^0 K^+ \pi^-$ is generated with a uniform amplitude or with equal and incoherent $K_0^*(1430)^- K^+$ and $\bar{K}_0^*(1430)^0 K^0$ contributions. The GAMGAM generator is also used to generate $\gamma\gamma \rightarrow \eta_c \rightarrow K_S^0 K^+ \pi^-$ events.

The analysis is performed with two data samples. The sample used to search for the process $\gamma\gamma \rightarrow X \rightarrow \eta_c \pi^+ \pi^-$ is referred to as the “main sample.” Properties of the η_c and its decay into $K_S^0 K^+ \pi^-$ are studied with a separate “control sample” of $\gamma\gamma \rightarrow \eta_c \rightarrow K_S^0 K^+ \pi^-$ events. For the main (control) sample, we select events that contain six (four) charged-particle tracks.

For both samples, charged kaon candidates are identified using likelihood values calculated from measurements of specific energy loss and information from a detector of internally reflected Cherenkov radiation. All other tracks are assumed to be pions. A K_S^0 candidate is reconstructed by fitting a $\pi^+ \pi^-$ pair to a common vertex, with invariant mass in the range $0.491 < m(\pi^+ \pi^-) < 0.503 \text{ GeV}/c^2$. A kinematic fit is performed, constraining $m(\pi^+ \pi^-)$ to the nominal K_S^0 mass [22]. An $\eta_c \rightarrow K_S^0 K^+ \pi^-$ decay candidate is reconstructed by combining a K_S^0 candidate with a K^+ and a π^- and requiring the resulting invariant mass to lie in the range $2.77 < m(K_S^0 K^+ \pi^-) < 3.22 \text{ GeV}/c^2$. In the main sample, the decay $X \rightarrow \eta_c \pi^+ \pi^-$ is reconstructed by combining an η_c candidate with the remaining two tracks in the event. A kinematic fit is applied, requiring the X -candidate decay vertex to be consistent with the e^+e^- interaction region. The angle $\alpha_{K_S^0}$ between the K_S^0 momentum vector and the line connecting the η_c and the K_S^0 decay vertices is required to satisfy $\cos \alpha_{K_S^0} > 0.99$.

In the control sample, we require the polar angle (the angle with respect to the beam axis) θ_{η_c} of the η_c candidate to satisfy $|\cos \theta_{\eta_c}| > 0.99$ and the transverse momentum of the η_c candidate to satisfy $p_{\eta_c}^T < 0.5 \text{ GeV}/c$, both in the center-of-mass (c.m.) frame of the e^+e^- system. The extra energy in the event, defined as the total energy in calorimeter clusters not associated with the identified tracks, is required to satisfy $E_{\text{ex}} < 0.5 \text{ GeV}$ in the c.m. frame. The $m(K_S^0 K^+ \pi^-)$ distribution of the selected control-sample events, shown in Fig. 1(a), exhibits clear η_c and J/ψ peaks, with the J/ψ produced in ISR events. In the main sample, continuum background is strongly suppressed with the requirements $|\cos \theta_X| > 0.85$, $p_X^T < 1.5 \text{ GeV}/c$, and $E_{\text{ex}} < 0.8 \text{ GeV}$, where $\cos \theta_X$ and p_X^T are the polar angle and transverse momentum of the X candidate. In addition, the total visible energy in the event, obtained from all charged tracks and calorimeter clusters, is required to satisfy $E_{\text{vis}} < 10 \text{ GeV}$ in the laboratory frame.

In the main sample, we suppress QED background by requiring $R2 < 0.7$, where $R2$ is the ratio of the second and zeroth Fox-Wolfram moments [23]. We suppress background due to ISR events with a requirement on the missing mass squared $m_{\text{miss}}^2 \equiv (p_{e^+e^-} - p_X)^2 > 10 \text{ GeV}^2/c^4$,

where $p_{e^+e^-}$ (p_X) is the total 4-momentum of the beam particles (X candidate).

Additional background suppression in the main sample is obtained by using the Dalitz plot for the η_c candidates. The Dalitz plot is shown in Fig. 1(b) for control-sample events in the η_c peak region $2.94 < m(K_S^0 K^+ \pi^-) < 3.02$ GeV/ c^2 , and in Figs. 1(c) and 1(d) for main-sample events in the lower and upper $m(K_S^0 K^+ \pi^-)$ sidebands $2.8 < m(K_S^0 K^+ \pi^-) < 2.9$ GeV/ c^2 and $3.05 < m(K_S^0 K^+ \pi^-) < 3.2$ GeV/ c^2 , respectively. These distributions indicate that true $\eta_c \rightarrow K_S^0 K^+ \pi^-$ decays often proceed via intermediate $K_0^*(1430)$ states, while background events contain $K^*(892)$ decays and random combinations. Taking advantage of this difference to suppress non- η_c background in the main sample, we require $|m^2(K_S^0 \pi^-) - M_{K_0^*(1430)}^2| < 0.5$ GeV $^2/c^4$ or $|m^2(K^+ \pi^-) - M_{K_0^*(1430)}^2| < 0.5$ GeV $^2/c^4$, and exclude events that satisfy $|m^2(K_S^0 \pi^-) - M_{K^*(892)}^2| < 0.35$ GeV $^2/c^4$ or $|m^2(K^+ \pi^-) - M_{K^*(892)}^2| < 0.2$ GeV $^2/c^4$, where M_R is the peak mass of resonance R [22]. The Dalitz-plot region selected by these criteria is enclosed within the solid lines in Figs. 1(b)–1(d). The criteria are the result of maximizing $\varepsilon_{\eta_c}^{DP} / \sqrt{\varepsilon_{SB}^{DP}}$, where $\varepsilon_{\eta_c}^{DP} = (63.5 \pm 3.2)\%$ is the efficiency of the Dalitz-plot

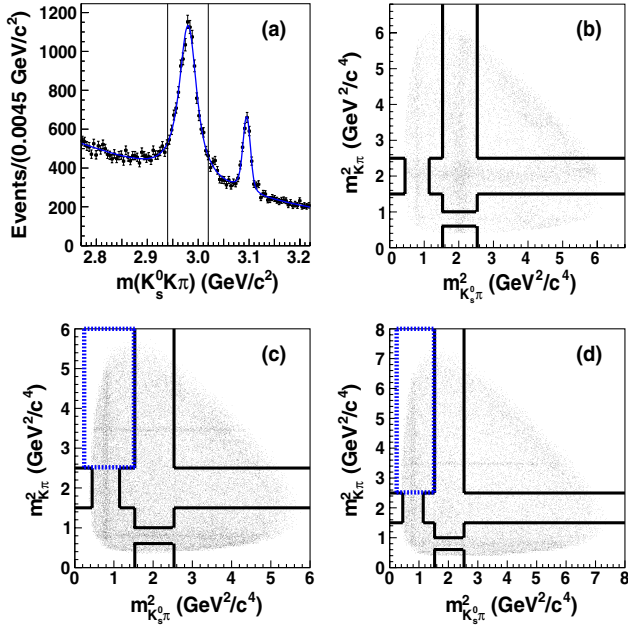


FIG. 1 (color online). (a) The $m(K_S^0 K^+ \pi^-)$ distribution for the control sample. The vertical lines indicate the η_c peak mass region. Also shown are the $K_S^0 K^+ \pi^-$ Dalitz plots for (b) control-sample events in the η_c peak mass region and for main-sample events in the (c) lower and (d) upper η_c mass sidebands. Solid black lines indicate the regions defined by the Dalitz-plot selection criteria. The dotted blue box in the upper left corner of (c) and (d) indicates the Dalitz-plot-sideband background region used for the neural-network training.

requirements for η_c decays, determined by fitting the $m(K_S^0 K^+ \pi^-)$ distribution of the control sample, and $\varepsilon_{SB}^{DP} = (30.74 \pm 0.21)\%$ is the corresponding efficiency for main-sample events in the $m(K_S^0 K^+ \pi^-)$ sidebands. The $\bar{D}^0 \rightarrow K^+ \pi^-$ band, evident in Fig. 1(c), becomes insignificant following a neural-network requirement, described below. Therefore, no explicit effort is made to remove this source of background.

Further background suppression is achieved by combining six variables into a neural-network discriminator. Two of the variables are E_{ex} and p_X^T . Each of the remaining four variables is a discrete output value of a kaon- or pion-identification algorithm applied to one of the charged-particle tracks that is not a daughter of the K_S^0 candidate. The distributions of the neural-network input variables are shown in Figs. 2(c) and 2(d). The neural network is trained with signal MC and main-sample background events in the Dalitz-plot sideband region $m^2(K^+ \pi^-) > 2.5$ GeV $^2/c^4$,

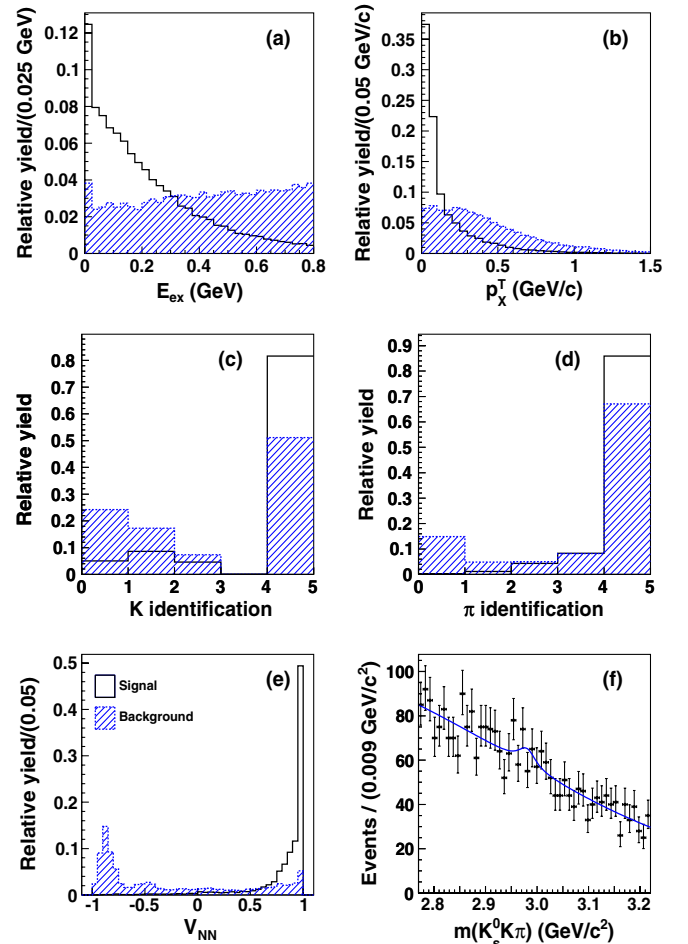


FIG. 2 (color online). Signal (unhatched) and background (hatched) distributions of the neural-network input variables (a) E_{ex} , (b) p_X^T , and the (c) kaon and (d) pion identification variables. (e) The distributions of the neural-network output variable V_{NN} . (f) The $m(K_S^0 K^+ \pi^-)$ distribution of the data with the step-1 fit function overlaid.

$m^2(K_S^0 \pi^-) < 1.5 \text{ GeV}^2/c^4$, indicated by the dashed boxes in Figs. 1(c) and 1(d). This region is chosen since it contains only $(3.40 \pm 0.66)\%$ of η_c decays in the control sample. We find only insignificant differences in the neural-network signal-to-background separation when using different signal samples or the mirror Dalitz-plot region $m^2(K^+ \pi^-) < 1.5 \text{ GeV}^2/c^4$, $m^2(K_S^0 \pi^-) > 2.5 \text{ GeV}^2/c^4$ for the background. The distributions of the output variable V_{NN} are shown in Fig. 2(e) for signal and background events. We find the optimal selection on this variable to be $V_{\text{NN}} > 0.84$. The signal efficiency of this selection is 72% for the $\eta_c(2S)$, and varies by up to 4%, depending on the X mass. The background efficiency is $(10.4 \pm 0.2)\%$ for the neural-network training region and $(7.4 \pm 0.2)\%$ for the mirror region.

We find 2863 main-sample events that satisfy all the selection criteria, with only about 700 events expected from non- $\gamma\gamma$ background MC. We conclude that the majority of the background is due to $\gamma\gamma$ events, for which we have no generic generator. More than one X candidate is reconstructed in 3.8% of the events. In these cases, we select the candidate for which $m^2(K_S^0 \pi^-)$ or $m^2(K^+ \pi^-)$ is closest to the $K_0^*(1430)$ peak.

In addition to these samples, an ISR-produced sample of $\psi(2S) \rightarrow J/\psi \pi^+ \pi^-$ events is used to evaluate a systematic uncertainty associated with the detector resolution. This sample is selected in the same way as the main sample, except that the neural-network and Dalitz-plot selections are not applied, the $K_S^0 K^+ \pi^-$ invariant mass is required to be between 3.0 and 3.2 GeV/c^2 , and m_{miss}^2 must be less than 1 GeV^2/c^4 .

We define four categories of events in the main sample: signal corresponds to $\gamma\gamma \rightarrow X \rightarrow \eta_c \pi^+ \pi^-$ events; combinatorial background (CB), which is by far the most copious background, arises from random combinations of final-state particles; events with a true $\eta_c \rightarrow K_S^0 K^+ \pi^-$ decay and two pions not originating from an X resonance decay are categorized as η_c -peaking background ($\eta_c\text{B}$); X -peaking background (XB) corresponds to decays $X \rightarrow K_S^0 K^+ \pi^- \pi^+ \pi^-$ that do not proceed through an intermediate η_c .

The extraction of the signal yield proceeds in two steps. In step 1, we determine the values of the $m(K_S^0 K^+ \pi^-)$ -distribution parameters of the combinatorial background from a one-dimensional fit to $m(K_S^0 K^+ \pi^-)$, without any restrictions on $m(K_S^0 K^+ \pi^- \pi^+ \pi^-)$.

In step 2, we extract the signal yield for each X resonance hypothesis from a two-dimensional fit to the $m(K_S^0 K^+ \pi^-)$ versus $m(K_S^0 K^+ \pi^- \pi^+ \pi^-)$ distribution for events in an $m(K_S^0 K^+ \pi^- \pi^+ \pi^-)$ window around the resonance peak. The fits use the unbinned, extended-maximum-likelihood method and are performed with the ROOFIT package [24].

From events in the $m(K_S^0 K^+ \pi^-)$ sidebands, we observe that all correlation between the $m(K_S^0 K^+ \pi^- \pi^+ \pi^-)$ and

$m(K_S^0 K^+ \pi^-)$ distributions for the combinatorial background is accounted for by the phase space $\Phi(m(K_S^0 K^+ \pi^-), m(K_S^0 K^+ \pi^- \pi^+ \pi^-))$ of the three-body final state consisting of the π^+ , π^- , and the $(K_S^0 K^+ \pi^-)$ system. This is used to construct the probability-density function (PDF) [25] of the step-1 fit. This PDF is a function of $m(K_S^0 K^+ \pi^-)$ with $m(K_S^0 K^+ \pi^- \pi^+ \pi^-)$ as a conditional variable, and is given by

$$\mathcal{H}(m_3|m_5) = N_{\eta_c} \mathcal{H}_{\eta_c}(m_3) + N_{\eta_c} \mathcal{H}_{\eta_c}(m_3|m_5), \quad (1)$$

where N_{η_c} (N_{η_c}) is the number of events with (without) a true $\eta_c \rightarrow K_S^0 K^+ \pi^-$ decay. We have used the notation $m_3 \equiv m(K_S^0 K^+ \pi^-)$ and $m_5 \equiv m(K_S^0 K^+ \pi^- \pi^+ \pi^-)$ for brevity.

The PDF for non- η_c events in Eq. (1) is

$$\mathcal{H}_{\eta_c}(m_3|m_5) = \mathcal{P}_2(m_3; a_1, a_2, m_3^0) \Phi(m_3, m_5), \quad (2)$$

where $\mathcal{P}_2(m_3; a_1, a_2, m_3^0) = 1 + a_1(m_3 - m_3^0) + a_2(m_3 - m_3^0)^2$ is a second-order polynomial and m_3^0 is a constant offset set to 3.0 GeV/c^2 in order to reduce correlations between a_1 and a_2 . Determination of the coefficients a_1, a_2 is the main purpose of the step-1 fit. The PDF for η_c events in Eq. (1) is

$$\mathcal{H}_{\eta_c}(m_3) = \mathcal{W}(m_3; M_{\eta_c}, \Gamma_{\eta_c}, \vec{r}_{m_3}), \quad (3)$$

where \mathcal{W} is a relativistic Breit-Wigner function $[(\tilde{m}_3^2 - M_{\eta_c}^2)^2 + M_{\eta_c}^2 \Gamma_{\eta_c}^2]^{-1}$ convolved with a detector-resolution function $\mathcal{R}(m_3 - \tilde{m}_3; \vec{r}_{m_3})$ that depends on a set of parameters \vec{r}_{m_3} and the difference between the measured m_3 and the true invariant mass \tilde{m}_3 of the $K_S^0 K^+ \pi^-$ system. The resolution function is the sum of two Crystal Ball functions [26] with oppositely directed tails and common Gaussian-parameter values,

$$\mathcal{R}(x; \vec{r}_x) = (a_l + a_r) \exp[-(x - \bar{x})^2/(2\sigma^2)] + a_r T_r(x) + a_l T_l(x), \quad (4)$$

where

$$T_i(x) = \begin{cases} A_i [\sigma/(x'_i - \bar{x})]^{n_i}; & x > \bar{x} + \alpha_i \sigma \\ 0; & x \leq \bar{x} + \alpha_i \sigma \end{cases} \quad (5)$$

and we have defined $x'_i = x + \sigma(n_i/\alpha_i - \alpha_i)$, and $A_i = (n_i/\alpha_i)^{n_i} \times \exp(-\alpha_i^2/2)$. In Eq. (4), the subscripts r and l label the parameters of the right and left tails, respectively. The values of the resolution-function parameters $\vec{r}_{m_3} = (\bar{x}, \sigma, n_i, \alpha_i, a_i)$ are determined from fits to the signal MC samples.

In addition to a_1 and a_2 , the parameter determined in the step-1 fit are the yields N_{η_c} and N_{η_c} , and the mass M_{η_c} and width Γ_{η_c} of the η_c peak. In order to obtain M_{η_c} and Γ_{η_c} from the data, the step-1 fit is performed simultaneously on the main sample and the control sample. The PDF for the control sample is

$$\begin{aligned}\mathcal{H}'(m_3) = & N'_{J/\psi} \mathcal{W}(m_3; M_{J/\psi}, \Gamma_{J/\psi}, \vec{r}_{m_3}) \\ & + N'_{\eta_c} \mathcal{H}_{\eta_c}(m_3) + N'_{\text{bgd}} \mathcal{P}_2(m_3; a'_1, a'_2, m_3^0).\end{aligned}\quad (6)$$

Additional control-sample parameter values determined in the fit are the peak J/ψ mass $M_{J/\psi}$, the background parameters a'_1, a'_2 , and the $\eta_c, J/\psi$, and background event yields $N'_{\eta_c}, N'_{J/\psi}$, and N'_{bgd} .

The $m(K_S^0 K^+ \pi^-)$ distribution of the data and the step-1 PDF are shown in Fig. 2(f). The fitted parameter values are $a_1 = 1.24 \pm 0.19 (\text{GeV}/c^2)^{-1}$, $a_2 = 0.2 \pm 1.4 (\text{GeV}/c^2)^{-2}$, $N_{\eta_c} = 50 \pm 37$, $N'_{\eta_c} = 10350 \pm 300$, and $N'_{J/\psi} = 1877 \pm 90$. The large relative uncertainties for a_1 and a_2 are the result of the near linearity of the $m(K_S^0 K^+ \pi^-)$ distribution and the correlation between the two parameters, which is taken into account in the evaluation of systematic uncertainties. The η_c parameter values determined in the step-1 fit are $\Gamma_{\eta_c} = 31.7 \pm 1.5 \text{ MeV}/c^2$ and $M_{\eta_c} = 2.98285 \pm 0.00038 \text{ GeV}/c^2$, where the uncertainties are statistical only. These results are consistent with previous measurements [22].

The PDF for the step-2 fit is a linear combination of the PDFs of the four event types,

$$\mathcal{P} = N_{\text{sig}} \mathcal{P}_{\text{sig}} + N_{\text{CB}} \mathcal{P}_{\text{CB}} + N_{\eta_c \text{B}} \mathcal{P}_{\eta_c \text{B}} + N_{\text{XB}} \mathcal{P}_{\text{XB}}. \quad (7)$$

The signal PDF is a relativistic Breit-Wigner function convolved with the resolution function, for both m_3 and m_5 :

$$\mathcal{P}_{\text{sig}}(m_3, m_5) = \mathcal{H}_{\eta_c}(m_3) \mathcal{W}(m_5; M_X, \Gamma_X, \vec{r}_{m_5}), \quad (8)$$

where M_X and Γ_X are the known mass and width of the resonance of interest [1,22,27], and \vec{r}_{m_5} are the parameters of the $m(K_S^0 K^+ \pi^- \pi^+ \pi^-)$ resolution function $\mathcal{R}(m_5 - \hat{m}_5; \vec{r}_{m_5})$, obtained from a fit to signal MC. The combinatorial-background PDF is

$$\mathcal{P}_{\text{CB}}(m_3, m_5) = \mathcal{H}_{\eta_c}(m_3 | m_5) \mathcal{C}_2(m_5; b_1^{\text{CB}}, b_2^{\text{CB}}), \quad (9)$$

where $\mathcal{C}_2(m_5; b_1^{\text{CB}}, b_2^{\text{CB}})$ is a second-order Chebychev polynomial with first- (second-) order coefficients b_1^{CB} (b_2^{CB}). The η_c -peaking background PDF is

$$\mathcal{P}_{\eta_c \text{B}}(m_3, m_5) = \mathcal{H}_{\eta_c}(m_3) \mathcal{C}_1(m_5; b_1^{\eta_c \text{B}}), \quad (10)$$

where $\mathcal{C}_1(m_5; b_1^{\eta_c \text{B}})$ is a first-order Chebychev polynomial. The X -peaking background PDF is

$$\mathcal{P}_{\text{XB}}(m_3, m_5) = \mathcal{P}_1(m_3; c_1^{\text{XB}}, m_3^0) \mathcal{W}(m_5; M_X, \Gamma_X, \vec{r}_{m_5}), \quad (11)$$

where $\mathcal{P}_1(m_3; c_1^{\text{XB}}, m_3^0)$ is a first-order polynomial. The parameter values determined with the step-2 fit are the four yields of Eq. (7) and the background shape parameters $b_1^{\text{CB}}, b_2^{\text{CB}}, b_1^{\eta_c \text{B}}$, and c_1^{XB} .

The step-2 fit is performed four times in different $m(K_S^0 K^+ \pi^- \pi^+ \pi^-)$ windows, fitting for the (1) $\chi_{c2}(1P)$, (2) $\eta_c(2S)$, (3) $X(3872)$ and $X(3915)$, or (4) $X(3872)$ and $\chi_{c2}(2P)$ resonances. A simultaneous fit to the three resonances $X(3872)$, $X(3915)$, and $\chi_{c2}(2P)$ is observed to be unstable when tested with parametrized MC experiments, due to the large number of fit parameters, small signal, and large overlap of the $X(3915)$ and $\chi_{c2}(2P)$ line shapes. Therefore, we conduct fits (3) and (4) separately to test for the existence of a signal for either set of line shape parameters. The $m(K_S^0 K^+ \pi^- \pi^+ \pi^-)$ and $m(K_S^0 K^+ \pi^-)$ distributions and fit functions are shown in Fig. 3. The difference between the fit function of fit (3) and that of fit (4) is almost indistinguishable within the thickness of the curve in Fig. 3(f). The fitted signal yields are summarized in Table I.

No significant signal or peaking background is observed in any of the fits. However, a hint of X -peaking background

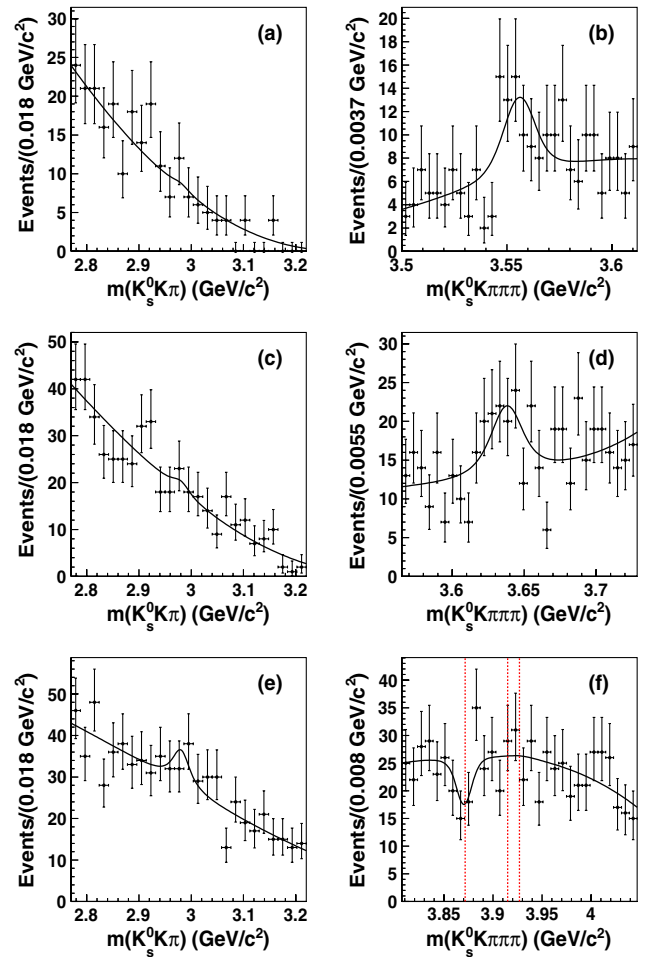


FIG. 3 (color online). Distributions of (a),(c),(e) $m(K_S^0 K^+ \pi^-)$ and (b),(d),(f) $m(K_S^0 K^+ \pi^- \pi^+ \pi^-)$ with the step-2 fit PDF overlaid for the fit regions of the (a),(b) $\chi_{c2}(1P)$, (c),(d) $\eta_c(2S)$, and (e),(f) $X(3872)$, $X(3915)$ and $\chi_{c2}(2P)$. The vertical dashed lines in (f) indicate the peak mass positions of the $X(3872)$, $X(3915)$, and $\chi_{c2}(2P)$ [22].

is visible in the $\chi_{c2}(1P)$ and $\eta_c(2S)$ fits of Figs. 3(b) and 3(d), with event yields of 33 ± 14 and 47 ± 24 , respectively, where the uncertainties are statistical only. This may be due to decays of $\chi_{c2}(1P)$ and $\eta_c(2S)$ into $K_S^0 K^+ \pi^- \pi^+ \pi^-$ [28], which are suppressed in this analysis by the $2.77 < m(K_S^0 K^+ \pi^-) < 3.22$ GeV/ c^2 requirement. Fits (3) and (4) yield insignificant X -peaking background, roughly canceling the negative signal yields. The results shown in Table I for the $X(3872)$ are obtained from fit (4). The $X(3872)$ yield from fit (3) is 1.6 events lower than in fit (4). Since no signal is observed for the $X(3915)$ or the $\chi_{c2}(2P)$, we obtain a conservative upper limit on the yield of the $X(3915)$ ($\chi_{c2}(2P)$) by fixing the $\chi_{c2}(2P)$ ($X(3915)$) yield to zero.

We estimate systematic uncertainties on the signal yields associated with the fit procedure by repeating the fits with the variations described below and adding the different uncertainties in quadrature. We account for uncertainties in the X mass and width values by varying them within their uncertainties [22]. This is the source of the largest signal-yield systematic uncertainty, except for the $\chi_{c2}(1P)$. The order of the polynomial in each PDF is varied to account for uncertainties due to background modeling. We vary the resolution-function Gaussian width [σ in Eq. (4)] by 2 MeV/ c^2 for the $m(K_S^0 K^+ \pi^-)$ PDF to account for a difference between the J/ψ mass resolution in MC and in the control sample, and by 0.9 MeV/ c^2 for the $m(K_S^0 K^+ \pi^- \pi^+ \pi^-)$ PDF due to a difference in the $\psi(2S)$ mass resolution between MC and data. An additional uncertainty is evaluated by using the sum of three Gaussians to define the resolution function. To address the possibility that correlations between the m_5 and m_3 distributions are not taken fully into account by the phase-space factor $\Phi(m_3, m_5)$ in Eq. (2), we replace the parameters a_i of Eq. (2) by $a_i(1 + a_i'' m_5)$. The values of the parameters a_i'' are found to be consistent with zero, and we conservatively use their uncertainties to evaluate the systematic uncertainty on the signal yield. The effect of not

accounting for phase-space correlations between m_3 and m_5 in the signal and η_c -peaking background PDFs is determined to be small compared to other systematic uncertainties, except for the $\chi_{c2}(1P)$, for which this uncertainty is dominant and equals 2.4 events. Statistical uncertainties from the step-1 fit are propagated to the step-2 fit, accounting for correlations among the parameters.

We test the entire fit procedure using parametrized MC experiments generated with the PDFs of Eqs. (1) and (7). A bias of up to two events on the signal yield is found and used as a correction that is accounted for in the values shown in Table I. A systematic uncertainty on this correction is evaluated by repeating this study after varying the generated signal yield by its statistical uncertainty in the data fit.

Since the dominant combinatorial background is distributed in phase space differently from signal and often contains additional final-state particles, interference between signal and background is expected to be relatively small. In addition, the small signal yields make the evaluation of such interference effects unreliable. Therefore, we do not attempt to account for possible interference.

We evaluate systematic uncertainties on track and K_S^0 reconstruction efficiencies, accounting for the momentum and angular distribution of signal tracks, as well as on the uncertainty of the Dalitz-plot requirement efficiency. A 2% systematic uncertainty is assigned due to differences between the distributions of the selection variables in the control sample and in $\gamma\gamma \rightarrow \eta_c \rightarrow K_S^0 K^+ \pi^-$ MC. Differences between the data and MC distributions of the particle-identification variables are studied using a high-purity sample of $D^{*+} \rightarrow \pi^+ D^0$, $D^0 \rightarrow K^- \pi^+$ events, and found to have negligible impact on the efficiency.

The expected Dalitz-plot dependence of the $X \rightarrow \eta_c \pi^+ \pi^-$ decay is unknown. However, we account for uncertainties in this amplitude, which is uniform in our simulated signal samples, by weighting events according to $(m^2(\pi\pi) - 4M_\pi^2)^2$ [5], where $m^2(\pi\pi)$ is the squared dipion mass and M_π is the π^- mass. From the weighted sample, we extract an efficiency correction of up to 4.6%

TABLE I. Results of the step-2 fits. For each resonance X , we show the peak mass and width used in the PDF (from Refs. [1,22,27]); the mass range of the fit; the efficiency; the bias-corrected signal yield with statistical and systematic uncertainties; the product of the $\gamma\gamma \rightarrow X$ production cross section and $X \rightarrow \eta_c \pi^+ \pi^-$ branching fraction, and the 90% C.L. upper limit (UL) on this product; the product of the two-photon partial width $\Gamma_{\gamma\gamma}$ and the $X \rightarrow \eta_c \pi^+ \pi^-$ branching fraction, and the 90% C.L. upper limit on this product. For the $X(3872)$ and the $X(3915)$ we assume $J = 2$.

Resonance	M_X (MeV/ c^2)	Γ_X (MeV)	m_5 range (GeV/ c^2)	ε (%)	N_{sig}	$\sigma\mathcal{B}$ (fb)		$\Gamma_{\gamma\gamma}\mathcal{B}$ (eV)	
						Central value	UL	Central value	UL
$\chi_{c2}(1P)$	3556.20 ± 0.09	1.97 ± 0.11	3.500–3.612	3.60 ± 0.39	$10.2^{+7.7}_{-6.3} \pm 3.5$	$37^{+28}_{-23} \pm 15$	80	$7.2^{+5.5}_{-4.4} \pm 2.9$	15.7
$\eta_c(2S)$	3638.5 ± 1.7	13.4 ± 5.6	3.565–3.728	3.53 ± 0.35	$17^{+12}_{-11} \pm 3$	$61^{+44}_{-41} \pm 16$	123	$65^{+47}_{-44} \pm 18$	133
$X(3872)$	3871.57 ± 0.25	3.0 ± 2.1	3.807–4.047	3.92 ± 0.38	$-4.7^{+7.9}_{-6.9} \pm 2.8$	$-16^{+26}_{-23} \pm 10$	38	$-4.5^{+7.7}_{-6.7} \pm 2.9$	11.1
$X(3915)$	3915.0 ± 3.6	17.0 ± 10.4	3.807–4.047	3.79 ± 0.37	$-13^{+11}_{-11} \pm 7$	$-44^{+38}_{-38} \pm 25$	53	$-13^{+12}_{-12} \pm 8$	16
$\chi_{c2}(2P)$	3927.2 ± 2.6	24 ± 6	3.807–4.047	3.75 ± 0.36	$-15^{+14}_{-13} \pm 4$	$-53^{+49}_{-46} \pm 18$	60	$-16^{+15}_{-14} \pm 6$	18

(incorporated into the values in Table I) and a systematic uncertainty of the same magnitude.

Finally, we account for a 0.45% uncertainty on the integrated luminosity, for the uncertainties on the K_S^0 , η_c , and $\eta_c(2S)$ branching fractions [22], and for MC-statistical uncertainties.

The results are summarized in Table I. From the signal yield N_{sig} of each resonance, the integrated luminosity L , and the signal efficiency ε , we compute the product $\sigma\mathcal{B} = N_{\text{sig}}/(L\varepsilon)$ of the $e^+e^- \rightarrow Xe^+e^-$ production cross section and the $X \rightarrow \eta_c\pi^+\pi^-$ branching fraction. We also evaluate the results in terms of the product $\Gamma_{\gamma\gamma}\mathcal{B}$, where $\Gamma_{\gamma\gamma}$ is the two-photon width of the resonance, by utilizing the GAMGAM generator to determine the cross section as a function of $\Gamma_{\gamma\gamma}$. A 4% uncertainty is assigned to the GAMGAM calculation [2]. Since we find no significant signal for the X resonances, we calculate 90% confidence-level (C.L.) Bayesian upper limits on these quantities, assuming a Gaussian likelihood incorporating statistical and systematic uncertainties.

Using the efficiency-corrected yields for the χ_{c2} and $\eta_c(2S)$ from [1], we find the relative branching fractions

$$\frac{\mathcal{B}(\eta_c(2S) \rightarrow \eta_c\pi^+\pi^-)}{\mathcal{B}(\eta_c(2S) \rightarrow K_S^0K^+\pi^-)} = 4.9_{-3.3}^{+3.5} \pm 1.3 \pm 0.8, \quad (12)$$

$$\frac{\mathcal{B}(\chi_{c2}(1P) \rightarrow \eta_c\pi^+\pi^-)}{\mathcal{B}(\chi_{c2}(1P) \rightarrow K_S^0K^+\pi^-)} = 14.5_{-8.9}^{+10.9} \pm 7.3 \pm 2.5,$$

where the first uncertainty is statistical, the second is systematic, and the third is due to the uncertainty on $\mathcal{B}(\eta_c \rightarrow K_S^0K^+\pi^-)$ [22]. The 90% C.L. upper limits on the two ratios in Eq. (12) are 10.0 and 32.9, respectively. Using $\mathcal{B}(\eta_c(2S) \rightarrow K_S^0K^+\pi^-)$ and $\mathcal{B}(\chi_{c2}(1P) \rightarrow K_S^0K^+\pi^-)$ from

Ref. [22], we obtain the 90% C.L. upper limits $\mathcal{B}(\eta_c(2S) \rightarrow \eta_c\pi^+\pi^-) < 7.4\%$ and $\mathcal{B}(\chi_{c2}(1P) \rightarrow \eta_c\pi^+\pi^-) < 2.2\%$.

In summary, we report a study of the process $\gamma\gamma \rightarrow \eta_c\pi^+\pi^-$ and provide, for the first time, upper limits on the branching fractions of $\chi_{c2}(1P)$ and $\eta_c(2S)$ decays to $\eta_c\pi^+\pi^-$ relative to the branching fractions of the decays into $K_S^0K^+\pi^-$. We also report upper limits on the products $\sigma\mathcal{B}$ and $\Gamma_{\gamma\gamma}\mathcal{B}$ for the $\chi_{c2}(1P)$, $\eta_c(2S)$, $X(3872)$, $X(3915)$, and $\chi_{c2}(2P)$ resonances.

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